#### Quasi-stationary distributions and the decay parameter

Hanjun Zhang

Department of Mathematics, University of Queensland

hjz@maths.uq.edu.au



AUSTRALIAN RESEARCH COUNCIL Centre of Excellence for Mathematics and Statistics of Complex Systems

#### Introduction

#### Introduction

Quasi-stationary distributions

#### Introduction

- Quasi-stationary distributions
- The decay parameter

- Introduction
- Quasi-stationary distributions
- The decay parameter
- The Generalised Markov Branching Processes (GMBP)

- Introduction
- Quasi-stationary distributions
- The decay parameter
- The Generalised Markov Branching Processes (GMBP)
- Conclusion

- Introduction
- Quasi-stationary distributions
- The decay parameter
- The Generalised Markov Branching Processes (GMBP)
- Conclusion
- Further research

We consider a continuous-time Markov process  $(X_t)$  with a countable state space, taken here to be  $\mathbb{N}_+$ , and with a single absorbing state 0.

We consider a continuous-time Markov process  $(X_t)$  with a countable state space, taken here to be  $\mathbb{N}_+$ , and with a single absorbing state 0.

Let  $Q = (q_{ij})$  denote the generator, assumed here to be stable, conservative and regular.

We consider a continuous-time Markov process  $(X_t)$  with a countable state space, taken here to be  $\mathbb{N}_+$ , and with a single absorbing state 0.

Let  $Q = (q_{ij})$  denote the generator, assumed here to be stable, conservative and regular.

We denote by  $P = (p_{ij}(t))$  transition probabilities of the minimal process which here is the unique process with generator Q.

We consider a continuous-time Markov process  $(X_t)$  with a countable state space, taken here to be  $\mathbb{N}_+$ , and with a single absorbing state 0.

Let  $Q = (q_{ij})$  denote the generator, assumed here to be stable, conservative and regular.

We denote by  $P = (p_{ij}(t))$  transition probabilities of the minimal process which here is the unique process with generator Q.

Let  $\mathbf{P_i}(\cdot) = \mathbf{P_i}(\cdot \mid \mathbf{X_0} = \mathbf{i})$  and If  $\nu$  is a finite measure on  $\mathbb{N}$ , let  $\mathbf{P_{\nu}} = \sum \nu_i \mathbf{P_i}$ . Here and below any unqualified sum is taken over  $\mathbb{N}$ .

We consider a continuous-time Markov process  $(X_t)$  with a countable state space, taken here to be  $\mathbb{N}_+$ , and with a single absorbing state 0.

Let  $Q = (q_{ij})$  denote the generator, assumed here to be stable, conservative and regular.

We denote by  $P = (p_{ij}(t))$  transition probabilities of the minimal process which here is the unique process with generator Q.

Let  $\mathbf{P_i}(\cdot) = \mathbf{P_i}(\cdot \mid \mathbf{X_0} = \mathbf{i})$  and If  $\nu$  is a finite measure on  $\mathbb{N}$ , let  $\mathbf{P_{\nu}} = \sum \nu_i \mathbf{P_i}$ . Here and below any unqualified sum is taken over  $\mathbb{N}$ .

Finally, assume that  $\mathbb{N}$  is irreducible and that 0 is accessible from some (and hence from every) state in  $\mathbb{N}$ .

We further define

$$T = \inf\{t \ge 0 : X(t) = 0\}$$

the absorption (hitting) time at 0. We shall only the interested in processes for which  $E_i T < \infty$  for all  $i \ge 1$ .

A quasi-stationary distribution (qsd)  $M = (m_i)$  is a probability measure on  $\{1, 2, \dots\}$  with the property that, starting with  $M = (m_i)$ , the conditional distribution, given the event that at time *t* the process has not been absorbed, still  $M = (m_i)$ . That is,

$$\frac{\sum m_i \mathbf{P}_i(\mathbf{X}(\mathbf{t}) = \mathbf{j})}{\sum m_i \mathbf{P}_i(\mathbf{X}(\mathbf{t}) \neq \mathbf{0})} = m_j.$$
(1)

Quasi-stationary distributions for Markov processes and chains have been studied by several authors. Vere-Jones (1962), Seneta and Vere-Jones (1996) and Kingman (1963) studied the case of a general denumerable state space.

Quasi-stationary distributions for Markov processes and chains have been studied by several authors. Vere-Jones (1962), Seneta and Vere-Jones (1996) and Kingman (1963) studied the case of a general denumerable state space.

Actually, there are a great deal of papers (almost over 300 papers ) dealing with the **qsd**s

We now distinguish several related notions; Anderson (1991), Chapter 5 is a good general reference.

We now distinguish several related notions; Anderson (1991), Chapter 5 is a good general reference.

Given  $\mu \ge 0$  we call a measure M on  $\mathbb{N}$  a  $\mu$ -invariant measure for Q if for each  $j \ge 1$ ,

$$\sum m_i q_{ij} = -\mu m_j, \tag{4}$$

We now distinguish several related notions; Anderson (1991), Chapter 5 is a good general reference.

Given  $\mu \ge 0$  we call a measure M on  $\mathbb{N}$  a  $\mu$ -invariant measure for Q if for each  $j \ge 1$ ,

$$\sum m_i q_{ij} = -\mu m_j, \tag{6}$$

and if for all t > 0,

$$\sum m_i p_{ij}(t) = e^{-\mu t} m_j,\tag{7}$$

it is called  $\mu$ -invariant on  $\{1, 2, \dots\}$  for P.

M.G. Nair and P.K.Pollett (1993)show that if M is probability distribution on  $\{1, 2, \dots\}$ . Then M is a quasi-stationary distribution on  $\{1, 2, \dots\}$  for P if and only if, for some  $\mu > 0$ , M is  $\mu$ -invariant on  $\{1, 2, \dots\}$  for P.

We call  $M = (m_j) \nu$ - the limit conditional distribution ( $\nu$ -LCD) if  $\nu$  is a probability measure on  $\{1, 2, \dots\}$  and each  $j \ge 1$ 

$$m_j = \lim_{t \to \infty} \mathbf{P}_{\nu}(\mathbf{X}_t = \mathbf{j} \mid \mathbf{T} > \mathbf{t})$$
(8)

exists and is a probability measure on  $\{1, 2, \cdots\}$  .

We call  $M = (m_j) \nu$ - the limit conditional distribution ( $\nu$ -LCD) if  $\nu$  is a probability measure on  $\{1, 2, \dots\}$  and each  $j \ge 1$ 

$$m_j = \lim_{t \to \infty} \mathbf{P}_{\nu}(\mathbf{X}_t = \mathbf{j} \mid \mathbf{T} > \mathbf{t})$$
(9)

exists and is a probability measure on  $\{1, 2, \cdots\}$  .

Trivially, any qsd M is an M-LCD.

We call  $M = (m_j) \nu$ - the limit conditional distribution ( $\nu$ -LCD) if  $\nu$  is a probability measure on  $\{1, 2, \dots\}$  and each  $j \ge 1$ 

$$m_j = \lim_{t \to \infty} \mathbf{P}_{\nu}(\mathbf{X}_t = \mathbf{j} \mid \mathbf{T} > \mathbf{t})$$
(10)

exists and is a probability measure on  $\{1, 2, \cdots\}$  .

Trivially, any qsd M is an M-LCD.

The  $\nu$ -LCD is a **qsd** (Vere-Jones(1996)).

A complete treatment of the **qsd** problem for a given family of processes should accomplish two things:

A complete treatment of the **qsd** problem for a given family of processes should accomplish two things:

(i) determination of all qsd's; and

A complete treatment of the **qsd** problem for a given family of processes should accomplish two things:

(i) determination of all qsd's; and

(ii) solve the domian of attraction problem, namely, characterize all probability measure  $\nu$  such that a given qsd M is a  $\nu$ -LCD.

A complete treatment of the **qsd** problem for a given family of processes should accomplish two things:

(i) determination of all qsd's; and

(ii) solve the domian of attraction problem, namely, characterize all probability measure  $\nu$  such that a given qsd M is a  $\nu$ -LCD.

Although (i) has been addressed for several cases, details about (ii) are known only for finite Markov processes, and for the subcritical MBP.

We now discuss the existence of **qsd** for a general Markov Chain.

We now discuss the existence of **qsd** for a general Markov Chain.

P.A. Ferrari, H. Kesten, S. Martinez and P. Picco (1995) prove the following interesting result which makes no reference to this general theory. They make the following definition of *asymptotic remoteness* (AR) of the absorbing state: For each t > 0

$$\lim_{i \to \infty} \mathbf{P_i}(\mathbf{T} > \mathbf{t}) = \mathbf{1}.$$
 (12)

In other words,  $T \Rightarrow \infty$  as  $i \to \infty$ .

Assume that AR condition holds, Ferrari et al. prove that a **qsd** exists iff

$$\mathbf{E_i}(\mathbf{e}^{\epsilon \mathbf{T}}) < \infty$$
 (13)

for some  $\epsilon > 0$  and  $i \in \mathbb{N}$ .

Assume that AR condition holds, Ferrari et al. prove that a **qsd** exists iff

$$\mathbf{E_i}(\mathbf{e}^{\epsilon \mathbf{T}}) < \infty$$
 (14)

for some  $\epsilon > 0$  and  $i \in \mathbb{N}$ .

Indeed this condition is necessary with, or without, AR condition.

Assume that AR condition holds, Ferrari et al. prove that a qsd exists iff

$$\mathbf{E_i}(\mathbf{e}^{\epsilon \mathbf{T}}) < \infty$$
 (15)

for some  $\epsilon > 0$  and  $i \in \mathbb{N}$ .

Indeed this condition is necessary with, or without, AR condition.

T. G. Pakes (1994) investigates what happens in a number of examples when AR condition fails. In fact, he examine quite closely two examples which violate AR condition but which nevertheless can have a **qsd**, showing AR condition is far from being a necessary condition, though it seems essential for the proofs of Ferrari et al.'s theorem.

First we have the following **Proposition 1** The following statements are equivalent:

1. Equation (22) holds, that is,

$$\mathbf{E_i}(\mathbf{e}^{\epsilon \mathbf{T}}) < \infty$$

for some  $\epsilon > 0$  and  $i \in \mathbb{N}$ .

2. There is  $\lambda$  with  $0 < \lambda < \inf_{i \ge 1} q_i$  (here  $q_i \equiv -q_{ii}$ ) for which the system

$$\sum_{j \neq i} q_{ij} x_j \le (q_i - \lambda) x_i - 1, \quad i \ge 1, \quad x_0 = 0$$
(16)

has a finite non-negative solution.

Secondly we can obtain that the following condition

$$\lim_{i \to \infty} \mathbf{E_i T} = \infty \tag{17}$$

can substitute for the AR condition (that is, for each t > 0 $\lim_{i\to\infty} \mathbf{P_i}(\mathbf{T} > \mathbf{t}) = 1$ .) which preserves the main result of Ferrari et all (1995).

**Remarks:** 1. It is easy to prove that AR condition  $\implies$  (17). So condition (17) is weaker than AR condition .

**Remarks:** 1. It is easy to prove that AR condition  $\implies$  (17). So condition (17) is weaker than AR condition .

2. As we know, the mean extinction time  ${\bf E}_i{\bf T}$  is the minimal non-negative solution of the system

$$\sum_{j\geq 0} q_{ij} z_j = -1, \quad i \geq 1, \quad z_0 = 0.$$
 (19)

So condition (17) is easier to check than AR condition.
# **Quasi-stationary distributions**

Our main result is as follows **Theorem 1** Assume that Q is stable, conservative and regular, and that Q restricted to  $\{1, 2, \dots\}$  is irreducible. Assume further that (17) holds, that is

 $\lim_{i\to\infty}\mathbf{E_iT}=\infty$ 

and that  $P_i(T < \infty) = 1$  for some (and hence all) *i*. Then a necessary and sufficient condition for the existence of a qsd is that there is  $\lambda$  with  $0 < \lambda < \inf_{i \ge 1} q_i$  for which the system (16)(that is,  $\sum_{j \ne i} q_{ij}x_j \le (q_i - \lambda)x_i - 1, i \ge 1, x_0 = 0$ ) has a finite non-negative solution.

Suppose we have a *q*-matrix *Q* over *E*. Let *P* be an arbitrary *Q*-transition function. Suppose that  $E = \{0\} \cup C$ , where 0 is an absorbing state and  $C = \{1, 2, \dots\}$  is irreducible. The decay parameter  $\lambda$  is defined by

$$\lambda = \lim_{t \to \infty} -\frac{1}{t} \log P_{ij}(t).$$

Kingman showed that this limit exists and is the same for all  $i, j \in C$ , and that  $0 \le \lambda < \infty$ .

It is called the decay parameter because there exist constants  $M_{ij} > 0$  with  $M_{ii} = 1$  such that

$$P_{ij}(t) \le M_{ij}e^{-\lambda t}, \qquad i, j \in C.$$

Note, in particular, that  $P_{ii}(t) \leq e^{-\lambda t}$ .

It is called the decay parameter because there exist constants  $M_{ij} > 0$  with  $M_{ii} = 1$  such that

$$P_{ij}(t) \le M_{ij}e^{-\lambda t}, \qquad i, j \in C.$$

Note, in particular, that  $P_{ii}(t) \leq e^{-\lambda t}$ .

Why are we interested in the decay parameter?

It is called the decay parameter because there exist constants  $M_{ij} > 0$  with  $M_{ii} = 1$  such that

$$P_{ij}(t) \le M_{ij}e^{-\lambda t}, \qquad i, j \in C.$$

Note, in particular, that  $P_{ii}(t) \leq e^{-\lambda t}$ .

Why are we interested in the decay parameter?

 $\lambda = \sup \{ \alpha : P_{ij}(t) = O(\exp[-\alpha t]) \text{ as } t \to \infty \ \forall i, j \in C \}.$ 

It is called the decay parameter because there exist constants  $M_{ij} > 0$  with  $M_{ii} = 1$  such that

$$P_{ij}(t) \le M_{ij}e^{-\lambda t}, \qquad i, j \in C.$$

Note, in particular, that  $P_{ii}(t) \leq e^{-\lambda t}$ .

Why are we interested in the decay parameter?

 $\lambda = \sup \{ \alpha : P_{ij}(t) = O(\exp[-\alpha t]) \text{ as } t \to \infty \ \forall i, j \in C \}.$ 

If  $\mu>\lambda,$  there does not exist any  $\mu\text{-invariant}$  measure (Pollett 1986); in particularly, there does not exist any qsd if  $\lambda=0$  .

Questions ? Two obvious problems now arise in the context of the decay parameter, namely,

**Questions** ? Two obvious problems now arise in the context of the decay parameter, namely,

• to give criteria for decay parameter  $\lambda$  to be positive in terms of the rates  $(q_{ij})$ ;

**Questions** ? Two obvious problems now arise in the context of the decay parameter, namely,

- to give criteria for decay parameter  $\lambda$  to be positive in terms of the rates  $(q_{ij})$ ;
- to determine the value of  $\lambda$ , or at least bounds for  $\lambda$ , in terms of the rates  $(q_{ij})$ .

**Example 1** Markov branching process (MBP).

**Example 1** Markov branching process (MBP).

We shall adopt the usual notation (Anderson (1991))in prescribing MBP, that is, let  $p_k, k \ge 0$ , denote a sequence of non-negative numbers such that  $\sum_{k=0}^{\infty} p_k = 1$ , and let

$$p(s) = \sum_{k=0}^{\infty} p_k s^k, \quad 0 \le s \le 1,$$

denote the probability generating function of this sequence.

**Example 1** Markov branching process (MBP).

We shall adopt the usual notation (Anderson (1991))in prescribing MBP, that is, let  $p_k, k \ge 0$ , denote a sequence of non-negative numbers such that  $\sum_{k=0}^{\infty} p_k = 1$ , and let

$$p(s) = \sum_{k=0}^{\infty} p_k s^k, \ 0 \le s \le 1,$$

denote the probability generating function of this sequence.

$$m = p'(1) = \sum_{k=0}^{\infty} k p_k.$$

MBP with generator Q is given by

$$q_{ij} = \begin{cases} 0 & \text{if } j < i - 1 \\ -ia(1 - p_i) & \text{if } j = i \\ iap_{j-i+1} & \text{if } j \ge i - 1, j \neq i. \end{cases}$$

MBP with generator Q is given by

$$q_{ij} = \begin{cases} 0 & \text{if } j < i - 1 \\ -ia(1 - p_i) & \text{if } j = i \\ iap_{j-i+1} & \text{if } j \ge i - 1, j \neq i. \end{cases}$$

It is well known that  $P_i(T < \infty) = 1$  iff  $m \le 1$ .

MBP with generator Q is given by

$$q_{ij} = \begin{cases} 0 & \text{if } j < i - 1 \\ -ia(1 - p_i) & \text{if } j = i \\ iap_{j-i+1} & \text{if } j \ge i - 1, j \neq i. \end{cases}$$

It is well known that  $P_i(T < \infty) = 1$  iff  $m \le 1$ .

And if  $m \leq 1$ , then the decay parameter is  $\lambda = (1 - m)a$ .

**Example 2**: The birth and death process

#### **Example 2**: The birth and death process

We shall adopt the usual notation in prescribing birth rates  $\lambda_i > 0$   $(i \ge 1)$ , with  $\lambda_0 = 0$ , and death rates  $\mu_i > 0$   $(i \ge 1)$ . Now define by  $\pi_1 = 1$  and

$$\pi_n = \prod_{k=2}^n \frac{\lambda_{k-1}}{\mu_k}, \qquad n \ge 2.$$

We will assume the process is absorbed with probability 1, that is,

$$\sum_{n=1}^{\infty} \frac{1}{\pi_n \lambda_n} = \infty.$$
 (21)

In order to state our main results, we need the following notation:

$$Q_n = \left(\frac{1}{\pi_1 \mu_1} + \sum_{j=1}^{n-1} \frac{1}{\pi_j \lambda_j}\right) \sum_{j=n}^{\infty} \pi_j, \quad n \ge 1,$$

and

$$S_0 = \sup_{n \ge 1} Q_n.$$

Phil Pollett and Hanjun Zhang have obtained the following

**Theorem 2** 
$$(4S_0)^{-1} \le \lambda \le S_0^{-1}$$
.

And, hence,

 $\lambda > 0$  if and only if  $S_0 < \infty$ .

For a general Markov chain, we have the following

For a general Markov chain, we have the following

**Proposition 2** If

$$\mathbf{E_i}(\mathbf{e}^{\epsilon \mathbf{T}}) < \infty$$
 (23)

for some  $\epsilon > 0$  and  $i \in \mathbb{N}$ , then the decay parameter  $\lambda > 0$ .

For a general Markov chain, we have the following

**Proposition 2 If** 

$$\mathbf{E_i}(\mathbf{e}^{\epsilon \mathbf{T}}) < \infty$$
 (24)

for some  $\epsilon > 0$  and  $i \in \mathbb{N}$ , then the decay parameter  $\lambda > 0$ .

By using **Propsosition 1**, we get

**Theorem 3** If there is  $\lambda$  with  $0 < \lambda < \inf_{i \ge 1} q_i$  for which the system

$$\sum_{j \neq i} q_{ij} x_j \le (q_i - \lambda) x_i - 1, \quad i \ge 1, \quad x_0 = 0$$
(25)

has a finite non-negative solution, then the decay parameter  $\lambda > 0$ .

**Theorem 3** If there is  $\lambda$  with  $0 < \lambda < \inf_{i \ge 1} q_i$  for which the system

$$\sum_{j \neq i} q_{ij} x_j \le (q_i - \lambda) x_i - 1, \quad i \ge 1, \quad x_0 = 0$$
(26)

has a finite non-negative solution, then the decay parameter  $\lambda > 0$ .

I guess the above condition is necessary for the decay parameter  $\lambda > 0$ .

**Theorem 3** If there is  $\lambda$  with  $0 < \lambda < \inf_{i \ge 1} q_i$  for which the system

$$\sum_{j \neq i} q_{ij} x_j \le (q_i - \lambda) x_i - 1, \quad i \ge 1, \quad x_0 = 0$$
(27)

has a finite non-negative solution, then the decay parameter  $\lambda > 0$ .

I guess the above condition is necessary for the decay parameter  $\lambda > 0$ .

We have the following interesting result:

**Corollary** If  $\sup_{i>1} \mathbf{E}_i \mathbf{T} < \infty$ , then  $\lambda > 0$ .

R. R. Chen (1997) and Anyue Chen (2002) have discussed the Generalized Markov Branching Processes, the generator Q is given by

R. R. Chen (1997) and Anyue Chen (2002) have discussed the Generalized Markov Branching Processes, the generator Q is given by

$$q_{ij} = \begin{cases} 0 & \text{if } j < i - 1 \\ -i^{\nu}a(1 - p_i) & \text{if } j = i \\ i^{\nu}ap_{j-i+1} & \text{if } j \ge i - 1, j \neq i. \end{cases}$$

where, as above,  $p_k, k \ge 0$ , denote a sequence of non-negative numbers such that  $\sum_{k=0}^{\infty} p_k = 1$ .

And let

$$p(s) = \sum_{k=0}^{\infty} p_k s^k, \quad 0 \le s \le 1,$$

denote the probability generating function of this sequence.

And let

$$p(s) = \sum_{k=0}^{\infty} p_k s^k, \quad 0 \le s \le 1,$$

denote the probability generating function of this sequence.

$$m = p'(1) = \sum_{k=0}^{\infty} k p_k.$$

And let

$$p(s) = \sum_{k=0}^{\infty} p_k s^k, \quad 0 \le s \le 1,$$

denote the probability generating function of this sequence.

$$m = p'(1) = \sum_{k=0}^{\infty} k p_k.$$

It can be seen that the ordinary MBP corresponds to the special case of  $\nu = 1$ .

#### R.R.Chen obtained the following conclusions

(i) If  $\nu > 1$ , then Q is regular if and only if  $m \le 1$ .

R.R.Chen obtained the following conclusions

- (i) If  $\nu > 1$ , then Q is regular if and only if  $m \le 1$ .
- (ii) If  $\nu \leq 1$ , then it is regular if  $m < \infty$ .

R.R.Chen obtained the following conclusions

(i) If  $\nu > 1$ , then Q is regular if and only if  $m \le 1$ .

(ii) If  $\nu \leq 1$ , then it is regular if  $m < \infty$ .

(iii) Assume the given GMBP Q is regular. Then the extinction probability of the corresponding GMBP is 1 if and only if  $m \leq 1$ .

R.R.Chen obtained the following conclusions

(i) If  $\nu > 1$ , then Q is regular if and only if  $m \le 1$ .

(ii) If  $\nu \leq 1$ , then it is regular if  $m < \infty$ .

(iii) Assume the given GMBP Q is regular. Then the extinction probability of the corresponding GMBP is 1 if and only if  $m \leq 1$ .

Recall that a conservative Q is called regular if the Feller minimal Q-process is honest and thus there exists unique Q-process.

Anyue Chen (2002) obtained the following conclusion

If assume that the probability of eventual extinction is 1, i.e., assume  $m \leq 1$ . Then for all  $i \geq 1$ ,  $\mathbf{E_iT}$  are finite if and only if

$$\int_{0}^{1} \frac{1-y}{p(s)-s} (-\ln y)^{\nu-1} dy < \infty.$$
 (28)

Anyue Chen (2002) obtained the following conclusion

If assume that the probability of eventual extinction is 1, i.e., assume  $m \leq 1$ . Then for all  $i \geq 1$ ,  $\mathbf{E_iT}$  are finite if and only if

$$\int_0^1 \frac{1-y}{p(s)-s} (-\ln y)^{\nu-1} dy < \infty.$$
 (30)

Moreover, If (30) is true, then for all  $i \ge 1$ 

$$E_i T = \frac{1}{\Gamma(\nu)} \int_0^1 \frac{1-y}{a(p(s)-s)} (-\ln y)^{\nu-1} dy < \infty.$$
 (31)

Where  $\Gamma(\nu)$  is the gamma function.
We here talk about the positivity of the decay parameter and the existence of qsd. The following conclusions are obtained

We here talk about the positivity of the decay parameter and the existence of qsd. The following conclusions are obtained

**Theorem 3** (i) If m < 1, and  $\nu \ge 1$ , then the decay parameter  $\lambda > 0$ .

We here talk about the positivity of the decay parameter and the existence of qsd. The following conclusions are obtained

**Theorem 3** (i) If m < 1, and  $\nu \ge 1$ , then the decay parameter  $\lambda > 0$ .

(ii) If m = 1 and  $\nu \ge 2$ , then the decay parameter  $\lambda > 0$ .

We here talk about the positivity of the decay parameter and the existence of qsd. The following conclusions are obtained

**Theorem 3** (i) If m < 1, and  $\nu \ge 1$ , then the decay parameter  $\lambda > 0$ .

(ii) If m = 1 and  $\nu \ge 2$ , then the decay parameter  $\lambda > 0$ .

(iii) If m = 1,  $1 < \nu \le 2$  and  $\sum_{k=1}^{\infty} k^2 p_k < \infty$ , then there exists a qsd.

**Theorem 1** Assume that Q is stable, conservative and regular, and that Q restricted to  $\{1, 2, \dots\}$  is irreducible. Assume further that (17) holds, that is

$$\lim_{i \to \infty} E_i T = \infty$$

and that  $P_i(T < \infty) = 1$  for some (and hence all) *i*. Then a necessary and sufficient condition for the existence of a **qsd** is that there is  $\lambda$  with  $0 < \lambda < \inf_{i \ge 1} q_i$  for which the system (16), that is,

$$\sum_{j \neq i} q_{ij} x_j \le (q_i - \lambda) x_i - 1, \ i \ge 1, \ x_0 = 0$$

has a finite non-negative solution.

The decay parameter plays an important role in studying properties of Markov chains.

The decay parameter plays an important role in studying properties of Markov chains.

If there is  $\lambda$  with  $0 < \lambda < \inf_{i \ge 1} q_i$  for which the system

$$\sum_{j \neq i} q_{ij} x_j \le (q_i - \lambda) x_i - 1, \quad i \ge 1, \quad x_0 = 0$$
(33)

has a finite non-negative solution, then the decay parameter  $\lambda > 0$ .

The decay parameter plays an important role in studying properties of Markov chains.

If there is  $\lambda$  with  $0 < \lambda < \inf_{i \ge 1} q_i$  for which the system

$$\sum_{j \neq i} q_{ij} x_j \le (q_i - \lambda) x_i - 1, \quad i \ge 1, \quad x_0 = 0$$
(34)

has a finite non-negative solution, then the decay parameter  $\lambda > 0$ .

**Corollary** If  $\sup_{i>1} \mathbf{E}_i \mathbf{T} < \infty$ , then  $\lambda > 0$ .

For every birth-death process satisfying (20), that is,

$$\sum_{n=1}^{\infty} \frac{1}{\pi_n \lambda_n} = \infty.$$

we have

$$(4S_0)^{-1} \le \lambda \le S_0^{-1},$$

and hence

$$\lambda > 0$$
 if and only if  $S_0 < \infty$ .

#### **Further research**

We wish to prove  $\sup_{i\geq 1} \mathbf{E}_i \mathbf{T} < \infty$ , then there exists a qsd.

#### **Further research**

● We wish to prove  $\sup_{i\geq 1} \mathbf{E_iT} < \infty$ , then there exists a qsd.

We will obtain some formulae for the values of the decay parameter in general Markov processes.

I would like to acknowledge the support of the Australian Research Council Centre of Excellence for Mathematics and Statistics of Complex Systems.

Thanks are due to Phil Pollett, Ben Cairns and David Sirl.